

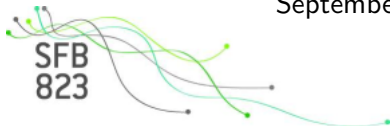
Asymptotic Theory for Copula Rank-Based Periodograms

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joint work with:

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Research School

Three Time Series Models, $X_t = Y_t/\text{Var}(Y_t)^{1/2}$

- QAR(1) process, Koenker and Xiao (2006),

$$Y_t = 0.1\Phi^{-1}(U_t) + 1.9(U_t - 0.5)Y_{t-1}$$

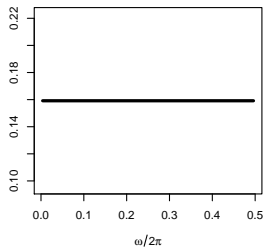
- (U_t) i.i.d. standard uniform random variables,
 - Φ the distribution function of the standard normal distribution.
- ARCH(1) process, Engle (1982),

$$Y_t = (1/1.9 + 0.9Y_{t-1}^2)^{1/2}\varepsilon_t$$

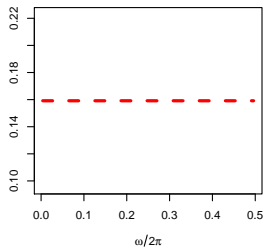
- (ε_t) standard normal white noise.
- Independent Gaussian white noise.

Spectral densities

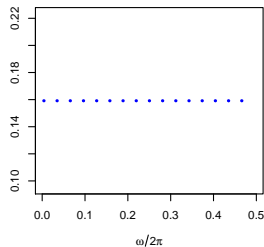
i.i.d.



QAR(1)



ARCH(1)

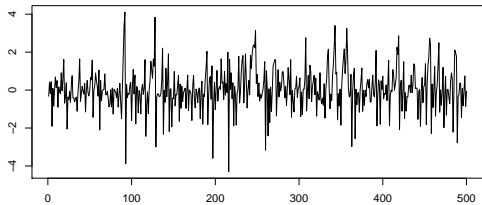


Spectral density:

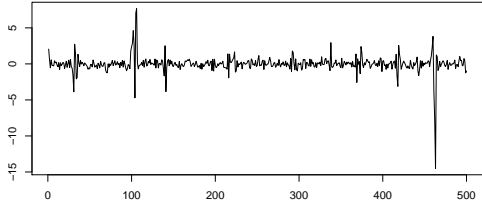
$$\frac{1}{2\pi} \sum_{k \in \mathbb{Z}} \text{Cov}(X_t, X_{t-k}) e^{-ik\omega}$$

$$X_t := Y_t / \text{Var}(Y_t)^{1/2}$$

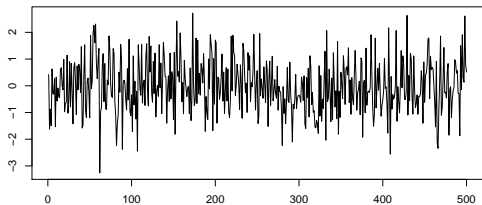
QAR(1)



ARCH(1)



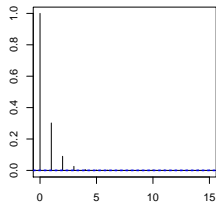
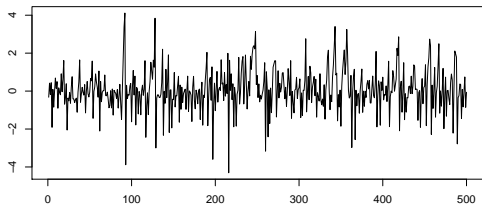
i.i.d. $\mathcal{N}(0,1)$



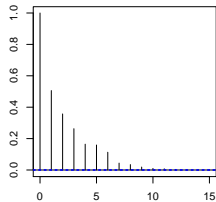
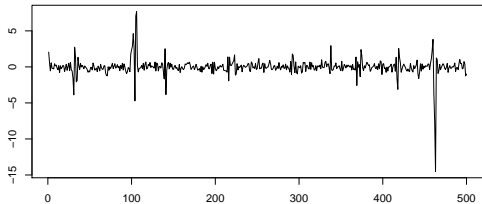
$$X_t := Y_t / \text{Var}(Y_t)^{1/2}$$

$$\text{Corr}(X_k^2, X_0^2)$$

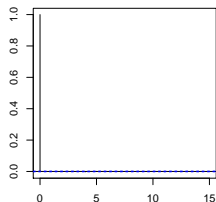
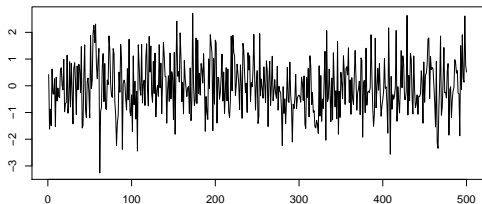
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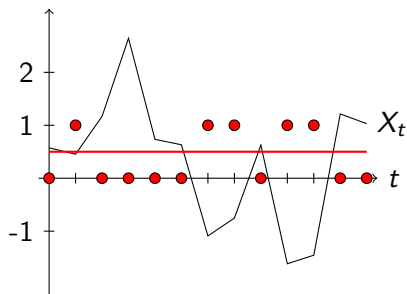


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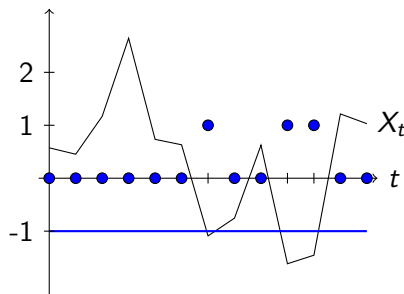
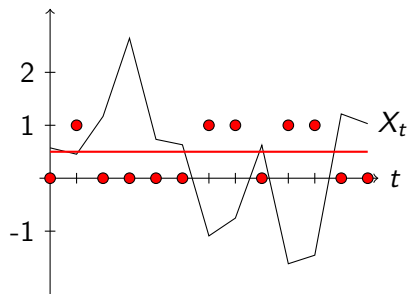
Clipped Process

$$(I\{X_t \leq 0.5\})_{t \in \mathbb{N}}$$



Clipped Processes

$(I\{X_t \leq 0.5\})_{t \in \mathbb{N}}$ and $(I\{X_t \leq -1\})_{t \in \mathbb{N}}$



A Quantile-Based Measure for Serial Dependence

(X_t) stationary

Traditionally: Auto-covariances of lag k :

$$\text{Cov}(X_t, X_{t-k})$$

Analysis of the spectral density

$$f(\omega) := \sum_{k=-\infty}^{\infty} \text{Cov}(X_t, X_{t-k}) e^{-ik\omega}$$

is analysis of the auto-covariances.

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A Quantile-Based Measure for Serial Dependence

(X_t) strictly stationary, F cdf of X_t

New approach: **Copula cross**-covariances of lag k :

$$\gamma_k(\tau_1, \tau_2) := \text{Cov}(I\{F(X_t) \leq \tau_1\}, I\{F(X_{t-k}) \leq \tau_2\})$$

Analysis of the **Copula** spectral density **kernel**

$$f^{\tau_1, \tau_2}(\omega) := \sum_{k=-\infty}^{\infty} \text{Cov}(I\{F(X_t) \leq \tau_1\}, I\{F(X_{t-k}) \leq \tau_2\})e^{-ik\omega}$$

is analysis of the **copula cross**-covariances.

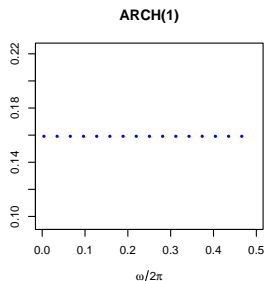
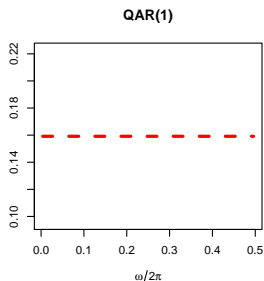
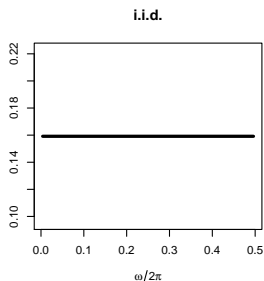
Covariance and Cross-covariance

- $\gamma_k(\tau_1, \tau_2)$ is cross-covariance of bivariate time series

$$(I\{F(X_t) \leq \tau_1\}, I\{F(X_t) \leq \tau_2\}),$$

- $\gamma_k(\tau_1, \tau_2)$ **always** exist (**no** assumptions about moments),
- $\gamma_k(\tau_1, \tau_2) = \mathbb{P}(F(X_t) \leq \tau_1, F(X_{t-k}) \leq \tau_2) - \tau_1\tau_2$
 \Rightarrow Copula: disentangling serial and marginal features
- **Invariance** of γ_k under continuous monotone transformation
- $\{\gamma_k(\tau_1, \tau_2) \mid \tau_1, \tau_2 \in (0, 1)\}$ and F **entirely** characterize the joint distribution of (X_t, X_{t-k}) ,
- ... if $\mathbb{E}X_t^2 < \infty$, then this includes the acf of $(X_t)_{t \in \mathbb{Z}}$,

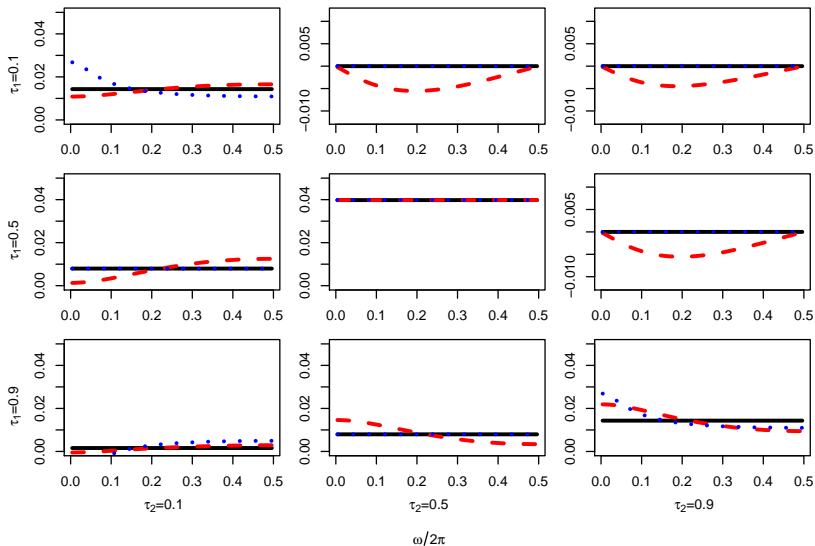
Spectral densities



Spectral density:

$$\frac{1}{2\pi} \sum_{k \in \mathbb{Z}} \text{Cov}(X_t, X_{t-k}) e^{-ik\omega}$$

Copula spectral density kernels



Modifying the (traditional) Periodogram

From X_0, \dots, X_{n-1} , for $\omega_j := \frac{2\pi j}{n} \in (0, \pi)$, $j \in \mathbb{Z}$, compute

$$I_n(\omega_j) := \frac{1}{2\pi n} d_n(\omega_j) d_n(-\omega_j),$$

where

$$d_n(\omega_j) := \sum_{t=0}^{n-1} X_t e^{-it\omega_j}$$

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$$d_{n,U}^T(\omega_j) := \sum_{t=0}^{n-1} I\{F(X_t) \leq \tau\} e^{-it\omega_j}$$

Modifying the (traditional) Periodogram

From X_0, \dots, X_{n-1} , for $\omega_j := \frac{2\pi j}{n} \in (0, \pi)$, $j \in \mathbb{Z}$, compute

$$I_{n,R}(\omega_j) := \frac{1}{2\pi n} d_{n,R}^{\tau_1}(\omega_j) d_{n,R}^{\tau_2}(-\omega_j),$$

where

$$d_{n,R}^{\tau}(\omega_j) := \sum_{t=0}^{n-1} I\{\hat{F}_n(X_t) \leq \tau\} e^{-it\omega_j}$$

$$\hat{F}_n(x) := \frac{1}{n} \sum_{t=0}^{n-1} I\{X_t \leq x\}$$

Asymptotic properties of $I_{n,R}^{\tau_1,\tau_2}(\omega)$

Theorem (K., Volgushev, Dette, Hallin (2015+))

Under suitable technical assumptions

$$I_{n,R}^{\tau_1,\tau_2}(\omega) \rightsquigarrow \mathbb{I}^{\tau_1,\tau_2}(\omega) \quad \text{in } \ell^\infty([0,1]^2) \quad \omega \in (0,\pi)$$

where $\mathbb{I}^{\tau_1,\tau_2}(\omega) := (2\pi)^{-1} \mathbb{D}^{\tau_1}(\omega) \overline{\mathbb{D}^{\tau_2}(\omega)}$, \mathbb{D} a centered, complex-valued Gaussian process with covariance structure

$$\text{Cov}(\mathbb{D}^{\tau_1}(\omega_1), \mathbb{D}^{\tau_2}(\omega_2)) = 2\pi f^{\tau_1,\tau_2}(\omega_1) I\{\omega_1 = \omega_2\}.$$

In particular: $\mathbb{I}^{\tau_1,\tau_2}(\omega_1), \mathbb{I}^{\tau_1,\tau_2}(\omega_2)$ independent for $\omega_1 \neq \omega_2$ and $\mathbb{E}[\mathbb{I}^{\tau_1,\tau_2}(\omega)] = f^{\tau_1,\tau_2}(\omega)$.

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Smoothing the CR periodogram kernel

- 'Asymptotic expectation' of $I_{n,R}^{\tau_1,\tau_2}(\omega)$ correct, no consistency
- Classical approach: smoothing

Definition (The smoothed CR periodogram kernel)

For kernel W and bandwidth b_n , let

$$\hat{f}_{n,R}^{\tau_1,\tau_2}(\omega) := \frac{2\pi}{n} \sum_{s=1}^{n-1} W_n(\omega - 2\pi s/n) I_{n,R}^{\tau_1,\tau_2}(2\pi s/n),$$

where $W_n(u) := \sum_{j=-\infty}^{\infty} b_n^{-1} W(b_n^{-1}[u + 2\pi j])$.

Asymptotic properties of $\hat{f}_{n,R}^{\tau_1, \tau_2}(\omega)$

Theorem (K., Volgushev, Dette, Hallin (2015+))

Under suitable assumptions, for any fixed $\omega \in (0, \pi)$

$$\mathbb{G}_n(\cdot, \cdot; \omega) := \sqrt{nb_n}(\hat{f}_{n,R}^{\tau_1, \tau_2}(\omega) - f^{\tau_1, \tau_2}(\omega) - B_n^{(k)}(\tau_1, \tau_2; \omega))_{\tau_1, \tau_2 \in [0,1]} \rightsquigarrow \mathbb{H}(\cdot, \cdot; \omega)$$

in $\ell^\infty([0, 1]^2)$, where $B_n^{(k)}(\tau_1, \tau_2; \omega) := \sum_{j=1}^k \frac{b_n^j}{j!} \int v^j W(v) dv \frac{d^j}{d\omega^j} f^{\tau_1, \tau_2}(\omega)$, and

$\mathbb{H}(\cdot, \cdot; \omega)$ is a centered, complex-valued Gaussian process with

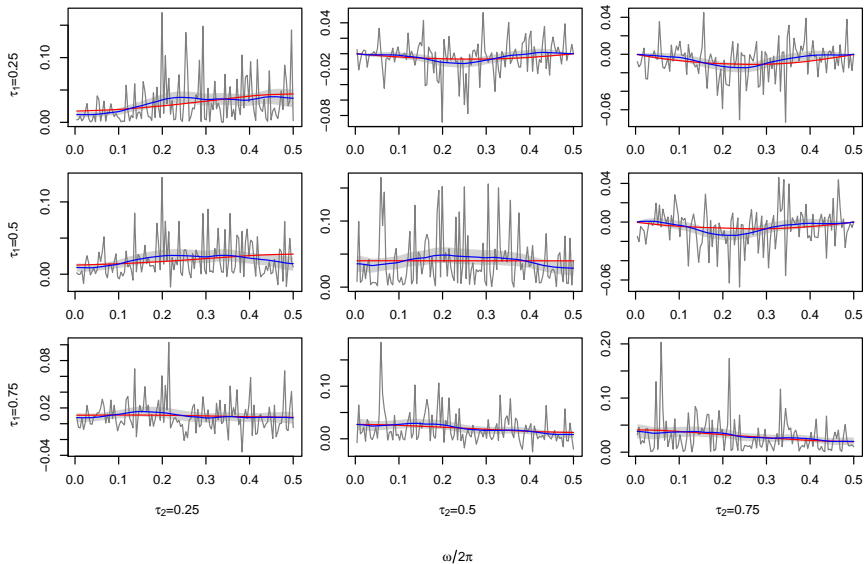
$$\text{Cov}(\mathbb{H}(x_1, y_1; \omega), \mathbb{H}(x_2, y_2; \omega)) = f^{x_1, y_1}(\omega) \overline{f^{x_2, y_2}(\omega)} \int W^2(u) du.$$

Moreover $\mathbb{H}(\omega) = \overline{\mathbb{H}(-\omega)} = \mathbb{H}(\omega + 2\pi)$ and $\{\mathbb{H}(\omega), \omega \in [0, \pi]\}$ is a family of independent processes. In particular, the weak convergence above holds jointly for finite, fixed collections of frequencies ω .

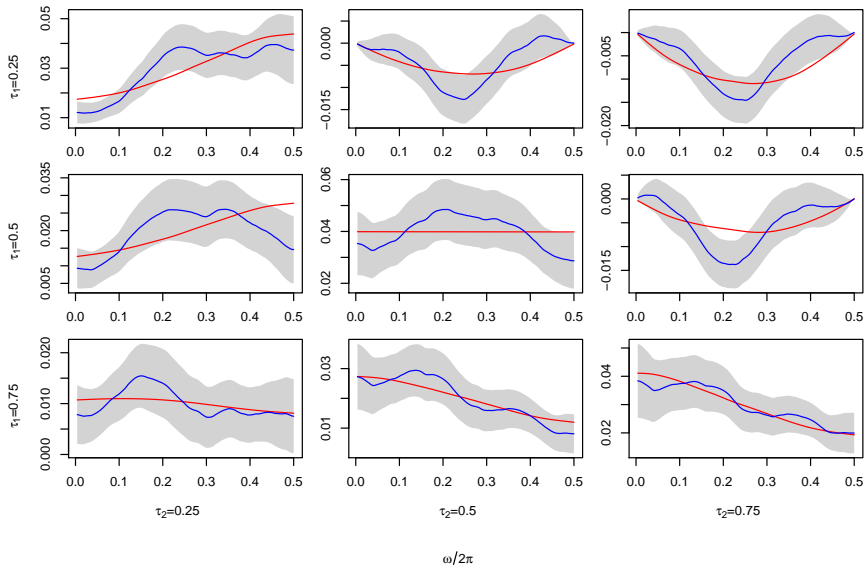
Example using the R-package quantspec

```
1 library(quantspec)
2
3 Y <- ts3(256)
4
5 levels <- c(0.25, 0.5, 0.75)
6 FF <- 2 * pi * (0:128) / 256
7 K <- length(levels)
8
9 wgt <- kernelWeight(W = W1, b = 0.1)
10
11 sPG.cl <- smoothedPG(Y, levels.1 = levels,
12   type = "clipped", weight = wgt)
13
14 sCSD <- quantileSD(N = 2^9, type = "copula", ts = ts3,
15   seed.init = 2581, levels.1 = levels, R = 1000)
16
17 plot(sPG.cl, plotPG = TRUE, qsd = sCSD, ratio = 1.7,
18   frequencies = FF[FF > 0], type.CIs = "naive.sd",
19   type.scaling = "individual")
```

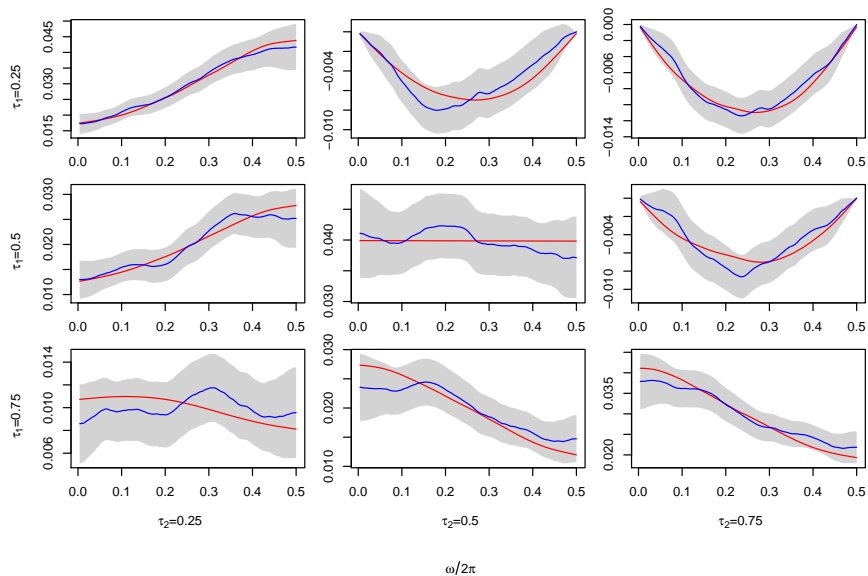
Plot of the SmoothedPG in the example, $n = 256$



Plot of the SmoothedPG in the example, $n = 256$



Plot of the SmoothedPG in the example, $n = 1.024$



“Model-free” and “nonlinear” spectral analysis

Quantile-based measures of serial dependence:

- Separation of serial dependencies and marginal features,
- Invariance under monotone transformations.

Quantile-based periodograms:

- inherit many of the properties of the ordinary periodogram,
- Robustness can be expected,
- Analysis of pair-copulae, not simply covariances,
- Weak convergence in $\ell^\infty([0, 1]^2)$,
- no linearity, distributional, nor even moment assumptions required.

Much work remains on the Research Agenda

- Tests based on the CR periodogram kernel,
- Estimation of higher-order spectra,
- Estimation of integrated spectra,
- Bootstrap,
- Locally stationary processes,
- ...

References

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- Kley, T. (2015). `quantspec`: Quantile-based Spectral Analysis Functions. R package version 1.0-3. Available on <http://cran.r-project.org/web/packages/quantspec/index.html>.