Valid confidence intervals for post-model-selection predictors

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Full linear model

(Full) linear model

$$Y = X\beta + U$$

- **Y** of size $n \times 1$: observation vector
- **X** of size $n \times p$: design matrix
- lacksquare eta of size $p \times 1$: regression coefficients
- $\mathbf{U} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_n)$
- p < n</p>

 \Longrightarrow Working distribution $P_{n,\beta,\sigma}$

Least square estimator:

$$\hat{\boldsymbol{\beta}} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{Y}$$

Standard variance estimator:

$$\hat{\sigma}^2 = \frac{1}{n-p} ||\mathbf{Y} - \mathbf{X}\hat{\boldsymbol{\beta}}||^2$$



Linear submodels

Linear submodels

Subsets $M \subset \{1,...,p\}$ of the columns of X. Give

$$\mathbf{Y} = \mathbf{X}[M]\mathbf{v} + \mathbf{U}$$

- M of cardinality m
- **X**[M] of size $n \times m$: only the columns of **X** that are in M
- \mathbf{v} of size $m \times 1$: needs to be defined/estimated to give the best representation of the full linear model

Non-standard regression coefficient vector

$$\beta_M^{(n)} = \underset{\mathbf{v}}{\operatorname{argmin}} ||\mathbf{X}\beta - \mathbf{X}[M]\mathbf{v}||$$
$$\beta_M^{(n)} = \beta[M] + (\mathbf{X}'[M]\mathbf{X}[M])^{-1} \mathbf{X}'[M]\mathbf{X}[M^c]\beta[M^c],$$

lacksquare eta[M] of size $m \times 1$: components of eta in M

Restricted least square estimator

$$\hat{\boldsymbol{\beta}}_{M} = \left(\boldsymbol{X}'[M]\boldsymbol{X}[M]\right)^{-1}\boldsymbol{X}'[M]\boldsymbol{Y}_{\square \rightarrow \square} + \mathbb{R} \rightarrow \mathbb{$$

The non-standard target of Berk et al.

Model selection procedure

Data-driven selection of the model with $\hat{M}(\mathbf{Y}) = \hat{M}$

Ex.: sequential testing, AIC, BIC, LASSO

Berk et al., 2013, Annals of Statistics consider the non-standard target

$$\boldsymbol{\beta}_{\hat{M}}^{(n)}$$

as their target for confidence intervals

Comments:

- Model selector \hat{M} is "imposed"
- Objective : best coefficients in this imposed model
- Random target

Predictors

Let \mathbf{x}_0 be a fixed $p \times 1$ vector and consider

$$y_0 = \boldsymbol{x}_0' \boldsymbol{\beta} + u_0$$

$$\mathbf{u}_0 \sim \mathcal{N}(\mathbf{0}, \sigma^2)$$

We consider the design-dependent non-standard target

$$\boldsymbol{x}_0'[\hat{M}]\boldsymbol{\beta}_{\hat{M}}^{(n)}$$

Optimality property: when x_0 is random and follows the empirical distribution given by the lines of X:

$$\mathbb{E}_{n,\boldsymbol{\beta},\sigma}\left(\left[y_0-\boldsymbol{x}_0'[\hat{M}]\boldsymbol{\beta}_{\hat{M}}^{(n)}\right]^2\right)\leq \mathbb{E}_{n,\boldsymbol{\beta},\sigma}\left(\left[y_0-\boldsymbol{x}_0'[\hat{M}]\boldsymbol{\nu}(\boldsymbol{Y})\right]^2\right),$$

for any function $v(Y) \in \mathbb{R}^{|\hat{M}|}$.



Confidence intervals

Let a nominal level $1 - \alpha \in (0, 1)$ be fixed

We consider confidence intervals for $\mathbf{x}_0'[\hat{M}]\beta_{\hat{M}}^{(n)}$ of the form

$$CI = \mathbf{x}_0'[\hat{M}]\hat{\boldsymbol{\beta}}_{\hat{M}} \pm K||\mathbf{s}_{\hat{M}}||\hat{\boldsymbol{\sigma}},$$

with

$$\mathbf{s}_{M}' = \mathbf{x}_{0}'[M] \left(\mathbf{X}'[M] \mathbf{X}[M] \right)^{-1} \mathbf{X}'[M]$$

Interpretation

- "Constant" K does not depend on Y (but on X, \mathbf{x}_0, \hat{M})
- \blacksquare For fixed M,

$$\mathbf{x}_0'[M]\hat{\boldsymbol{\beta}}_M - \mathbf{x}_0'[M]\boldsymbol{\beta}_M^{(n)} \sim \mathcal{N}(0, ||\mathbf{s}_M||\sigma^2)$$

- Thus, $K_{naive} = q_{S,n-p,1-\alpha/2}$ (Student quantile) is valid when M is deterministic
- When \hat{M} is random, K needs to be larger (e.g. Leeb et al. 2015, Statistical Science)
- \Longrightarrow Main issue : choosing K



Observe that

$$\mathbf{x}_0'[\hat{M}]\hat{eta}_{\hat{M}} - \mathbf{x}_0'[\hat{M}]eta_{\hat{M}}^{(n)} = \mathbf{s}_{\hat{M}}'(\mathbf{Y} - \mathbf{X}eta)$$

Then, we have

$$\left|\frac{\mathbf{s}_{\hat{M}}^{\prime}}{||\mathbf{s}_{\hat{M}}^{\prime}||\hat{\sigma}}\left(Y-\boldsymbol{X}\boldsymbol{\beta}\right)\right| \leq \max_{M\subseteq\{1,\ldots,\rho\}}\left|\frac{\mathbf{s}_{M}^{\prime}}{||\mathbf{s}_{M}^{\prime}||\hat{\sigma}}\left(Y-\boldsymbol{X}\boldsymbol{\beta}\right)\right|$$

Distribution of the upper-bound does not depend on β , $\sigma \Longrightarrow \text{let } K_1$ be its $(1 - \alpha)$ quantile

The CI given by K_1 satisfies

$$\inf_{\boldsymbol{\beta} \in \mathbb{R}^p, \sigma > 0} P_{n,\boldsymbol{\beta},\sigma} \left(\mathbf{x}_0'[\hat{M}] \boldsymbol{\beta}_{\hat{M}}^{(n)} \in CI \right) \ge 1 - \alpha$$

→ Uniformly valid confidence interval



Construction of new confidence intervals

The constant K_1 depends on all the components of \mathbf{x}_0

It can happen that only $\mathbf{x}_0[\hat{M}]$ is observed

model selection for cost reason

We construct other constants (see the paper for details)

$$K_1 \leq K_2 \leq K_3 \leq K_4$$

(The CIs given by K_2 , K_3 , K_4 are hence universally valid)

Remark : The case where only $\mathbf{x}_0[\hat{M}]$ is observed motivates all the more the study of $\mathbf{x}_0'[\hat{M}]\beta_{\hat{M}}^{(n)}$ as opposed to $\mathbf{x}_0'\mathcal{B}$

Design-independent non-standard target

Issue : The target $\mathbf{x}_0'[\hat{M}]\beta_{\hat{M}}^{(n)}$ depends on \mathbf{X}

Issue is solved when lines of $\textbf{\textit{X}}$ and $\textbf{\textit{x}}_0'$ are realizations from the same distribution $\mathcal L$

Let, for $\mathbf{x}'\sim\mathcal{L}$, $\mathbf{\Sigma}=\mathbb{E}(\mathbf{x}\mathbf{x}')$. Then, define the design-independent non-standard target by

$$\mathbf{x}_0[\hat{M}]'\beta_{\hat{M}}^{(\star)} = \mathbf{x}_0[\hat{M}]'\beta[\hat{M}] + \mathbf{x}_0[\hat{M}]'\left(\mathbf{\Sigma}[\hat{M},\hat{M}]\right)^{-1}\mathbf{\Sigma}[\hat{M},\hat{M}^c]\beta[\hat{M}^c],$$

Then, we have for $\mathbf{x}_0 \sim \mathcal{L}$,

$$\mathbb{E}\left(\left[y_0 - \mathbf{x}_0'[\hat{M}]\beta_{\hat{M}}^{(\star)}\right]^2\right) \leq \mathbb{E}\left(\left[y_0 - \mathbf{x}_0'[\hat{M}]\mathbf{v}(\mathbf{Y})\right]^2\right),$$

for any function $\mathbf{v}(\mathbf{Y}) \in \mathbb{R}^{|\hat{M}|}$



Asymptotic coverage when p is fixed and $n \to \infty$

Observe that

$$\begin{split} \left(\mathbf{x}_0[\hat{M}]'\boldsymbol{\beta}_{\hat{M}}^{(\star)} - \mathbf{x}_0[\hat{M}]'\boldsymbol{\beta}_{\hat{M}}^{(n)}\right) = \\ \mathbf{x}_0'[\hat{M}] \left(\left(\mathbf{X}'[\hat{M}]\mathbf{X}[\hat{M}]\right)^{-1}\mathbf{X}'[\hat{M}]\mathbf{X}[\hat{M}^c] - \left(\mathbf{\Sigma}[\hat{M},\hat{M}]\right)^{-1}\mathbf{\Sigma}[\hat{M},\hat{M}^c]\right)\boldsymbol{\beta}[\hat{M}^c] \end{split}$$

Theorem

Assume that

$$\sqrt{n}\left[\left(\mathbf{X}'\mathbf{X}/n\right)-\mathbf{\Sigma}\right]=O_p(1)$$

and that for any M with |M| < p and for any $\delta > 0$,

$$\sup\left\{P_{n,\beta,\sigma}(\hat{M}=M|\boldsymbol{X}):\boldsymbol{\beta}\in\mathbb{R}^{p},\sigma>0,\left\|\boldsymbol{\beta}[\boldsymbol{M}^{c}]\right\|/\sigma\geq\delta\right\}=o_{p}(1)$$

Then, for CI obtained by K_1, K_2, K_3, K_4 ,

$$\inf_{\boldsymbol{\beta} \in \mathbb{R}^{\rho}, \sigma > 0} P_{n,\boldsymbol{\beta},\sigma} \left(\mathbf{x}_0'[\hat{\boldsymbol{M}}] \boldsymbol{\beta}_{\hat{\boldsymbol{M}}}^{(\star)} \in CI \middle| \mathbf{X} \right) \geq (1 - \alpha) + o_{\rho}(1)$$



Simulation study

For $\alpha = 0.05$ and p = 10 we evaluate

$$\inf_{\boldsymbol{\beta} \in \mathbb{R}^{p}, \sigma > 0} P_{n, \boldsymbol{\beta}, \sigma} \left(\left. \boldsymbol{x}_{0}^{\prime}[\hat{\boldsymbol{M}}] \boldsymbol{\beta}_{\hat{\boldsymbol{M}}}^{(n, \star)} \in \textit{CI} \right| \boldsymbol{\textit{X}} \right),$$

for one realization of X

Results:

n	model	target							
	selector	design-dependent				design-independent			
		$\mathbf{x}_0[\hat{M}]'oldsymbol{eta}_{\hat{M}}^{(n)}$				$\mathbf{x}_0[\hat{M}]'oldsymbol{eta}_{\hat{M}}^{(\star)}$			
		K _{naive}	K_1	ÏK₃	K_4	K _{naive}	K_1	̈́K₃	K_4
20	AIC	0.84	0.99	1.00	1.00	0.79	0.97	0.99	0.99
20	BIC	0.84	0.99	1.00	1.00	0.74	0.96	0.98	0.98
20	LASSO	0.90	1.00	1.00	1.00	0.18	0.48	0.61	0.61
100	AIC	0.87	0.99	1.00	1.00	0.88	0.99	1.00	1.00
100	BIC	0.88	0.99	1.00	1.00	0.87	0.99	1.00	1.00
100	LASSO	0.88	0.99	1.00	1.00	0.87	0.99	1.00	1.00

Conclusion

- It is known that in the classical case (estimation of β), it is difficult to construct valid post-model-selection confidence intervals
- Recently, alternative coefficients have been studied Berk et al. 2013
 - this removes some obstacles
 - but naive procedures still fail
- We extend the confidence intervals to prediction
 - exact coverage of the design-dependent target
 - asymptotic coverage of the design-independent target

The paper :

F. Bachoc, H. Leeb, B.M. Pötscher. Valid confidence intervals for post-model-selection predictors,

http://arxiv.org/abs/1412.4605

Thank you for your attention!

